Maternal age-specific risks for adverse birth weights according to gestational weight gain: a prospective cohort in Chinese women older than 30

Yidi Wang¹, Yunhui Gong², Yujie Xu³, Xiaoyu Wang³, Shufang Shan³, Guo Cheng^{3*} and Ben Zhang^{1*}

Abstract

Background It is unclear whether the effects of abnormal gestational weight gain (GWG) on birth outcomes are differently in women with different maternal ages. This study aimed to investigate maternal age-specific association between GWG and adverse birth weights in Chinese women older than 30.

Methods 19,854 mother-child dyads were selected from a prospective cohort study in Southwest China between 2019 and 2022. Logistic regression model was used to assess the association between GWG, which defined by the 2009 Institute of Medicine guidelines, and adverse birth weights including large- and small-for-gestational-age (LGA and SGA), stratified by maternal age (31–34 years and ≥ 35 years).

Results In both maternal age groups, excessive and insufficient GWG were associated with increased odds of LGA and SGA, respectively. After women were categorized by pre-pregnancy body mass index, the associations remained significant in women aged 31-34 years, whereas for women aged ≥ 35 years, the association between excessive GWG and the risk of LGA was only significant in normal weight and overweight/obese women, and the significant effect of insufficient GWG on the risk of SGA was only observed in underweight and overweight/obese women. Moreover, among overweight/obese women, the magnitude of the association between insufficient GWG and the risk of SGA was greater in those aged ≥ 35 years (31-34 years: OR 2.08, 95\% Cl 1.19-3.55; ≥ 35 years: OR 2.65, 95\% Cl 1.47-4.74), while the impact of excessive GWG on the risk of LGA was more pronounced in those aged 31-34 years: OR 2.18, 95\% Cl 1.68-2.88; ≥ 35 years: OR 1.71, 95% Cl 1.30-2.25).

Conclusions The stronger associations between abnormal GWG and adverse birth weights were mainly observed in women aged 31–34 years, and more attention should be paid to this age group.

Keywords Gestational weight gain, Older maternal age, Body mass index, Large-for-gestational-age, Small-for-gestational-age

*Correspondence: Guo Cheng gcheng@scu.edu.cn Ben Zhang ben.zhang@scu.edu.cn

Full list of author information is available at the end of the article



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Background

Excessive or insufficient gestational weight gain (GWG) has been associated with a wide range of adverse perinatal outcomes, including large-for-gestational-age (LGA) [1] and small-for-gestational-age (SGA) [2]. Many studies have reported that LGA and SGA are associated with early morbidity and mortality [3, 4] as well as an increased risk of chronic health conditions such as obesity [5], diabetes [5], and hypertension [6] later in life. Therefore, maintaining an appropriate weight during pregnancy is important. Many factors may affect GWG, among which maternal age receives much attention as a trend in delayed childbearing has been observed in many countries [7, 8]. Advanced maternal age (AMA) is commonly defined as pregnancy at the age of 35 and older [9]. In China, the fertility rate for women aged 35–39 years was 5.7‰ in 1995 and rose to 18.6‰ in 2015 [10]. Furthermore, the number of women aged \geq 35 years has increased from 8.5 to 13.5% after the announcement of the two-child policy [11]. Older women usually gain less weight during pregnancy [12], and they are at a higher risk for a range of adverse perinatal outcomes, given that women aged 35-39 years had a 1.31-fold increased risk for LGA compared to those aged 25–29 years [13], and the SGA risk was increased 1.46-fold in women older than 40 years compared with women aged < 35 years [14]. Considering these age-related differences, it is of interest to know whether the association between GWG and adverse birth weight still exists in women of AMA, however, such evidence is scarce.

Although the age \geq 35 years has been suggested to be the cut-off for increased risk of adverse perinatal outcomes, some studies have reported that the risks were evident in women as early as 30-34 years by comparing with those aged 20-29 years [13, 15]. The relationship between GWG and adverse birth weights has been examined in women with mean maternal age at 30-34 years [1, 16–20], but these studies were limited by either small sample size [1, 16, 18-20] or retrospective design [16, 17, 19], and most of them were conducted in Western countries [1, 17-20]; little is known about the impact of GWG in Chinese women older than 30. Over the last two decades, the fertility rate for Chinese women aged 30-34 has risen from 26.5‰ to 45.3‰ (1995-2015), whereas it declined for women aged below 30 [10]. The mean age at childbearing in China has increased by 3.3 years during the same period (25.2 years in 1995 to 28.5 years in 2015) [10]. With the relaxation of the child policy, the mean childbearing age has increased to 29.6 years in 2019-2020 [21] and may exceed 30 years in the near future. Considering the increasing trend in women who give birth over 30, a prospective cohort with a large sample size to explore the role of GWG on adverse birth weights in this older Chinese population but not yet considered as AMA and then compare with those of AMA may have important implications.

Using the large prospective cohort from Southwest China, this study thus aimed to investigate the agespecific association between GWG and adverse birth weights in women above 30.

Methods

Study population

This prospective cohort study was conducted in the three provinces of Southwest China (Sichuan, Yunnan, and Guizhou provinces) between 2019 and 2022, aiming to investigate the effect of maternal nutritional status concerning pre-pregnancy body mass index (BMI) and GWG, on maternal and neonatal health outcomes in older mothers. Using the method of multistage sampling, 18 public hospitals and community health care centers were randomly selected from 12 urban and rural areas (1–2 hospitals in each area) in the three provinces. Pregnant women at their first prenatal visit (9⁺⁰-11⁺⁶ gestational week) were invited to participate if they were older than 30 years of age, singleton pregnancy, and had lived in their current residence for at least one year. Women with preexisting conditions such as diabetes, hypertension or other major diseases were excluded. Each pregnant women was scheduled to visit obstetricians (every 4 weeks through 25 weeks of gestation, every 2 weeks from 26 to 33 weeks of gestation, and weekly thereafter until birth) for medical examination and anthropometric measurements. Study participants were followed up at each prenatal visit until they gave birth. The study was approved by the Ethics Committee of the Sichuan University. Written informed consent was obtained from all participants.

During 2019 to 2022, there were 21,521 women recruited in this study. Pregnant women without information on pre-pregnancy or end pregnancy weight, neonatal outcomes or having stillbirth were excluded, resulting in a total of 19,854 mother-child dyads in the analysis of total GWG. Additional measure of GWG rate in the second and third trimesters (n 12,801) was used to assess the robustness of the results. The flow chart of the participants is shown in Figure S1.

Maternal measurements

Basic information including socio-demographic characteristics, medical history, family history of chronic diseases, and pre-pregnancy body weight was collected through a self-administrated questionnaire as well as women's height was measured during the first prenatal visit. Pre-pregnancy BMI was calculated [weight (kg) / height (m²)], and was further categorized into three groups according to World Health Organization Asian BMI classification [22]: underweight (<18.5 kg/m²), normal weight (18.5–22.9 kg/m²), and overweight/obese (\geq 23 kg/m²).

Anthropometric measures including maternal body weight during pregnancy was measured at each prenatal visit, and information on pregnancy complications was obtained from the Medical Birth Registry. GWG was calculated as the difference between the latest weight before delivery and pre-pregnancy weight. Women were categorized as less than, within, or greater than the 2009 Institute of Medicine (IOM)'s pre-pregnancy BMI-specific GWG recommendation [23].

To assess the robustness of the results, the rate of GWG in the second and third trimesters was estimated as [the difference between the first weight recorded in the second trimester $(12^{+1}-15^{+6}$ gestational week) and the last weight recorded before delivery] / (gestational age –13)] [23]. The rate of GWG in the second and third trimesters was classified following the 2009 IOM recommendation [23].

Birth weights

Birth weight was measured by trained medical workers and were retrieved from the Medical Birth Registry. LGA and SGA were defined as birth weight above the 90th percentile or below the 10th percentile, respectively, after adjusting for gender and gestational age according to the Chinese neonatal birth weight curve [24]. Gestational age was estimated based on the women's last menstrual period and was confirmed with ultrasound examination.

Statistical analysis

Continuous variables were presented as median with 25% and 75% quartiles as they were not normally distributed, and categorical variables were reported as frequency and proportion. The Chi-square test and Wilcoxon test were used to compare maternal and neonatal characteristics between subgroups. Logistic regression was conducted to estimate the unadjusted odd ratios (OR), adjusted OR, and 95% confidence intervals (CI) of adverse birth weights across GWG categories. GWG within the IOM recommendation was used as a reference group. Variables that were risk factors for adverse birth weights based on literature [14, 25-27] and were statistically different according to maternal GWG (Table S1) were identified as potential confounders: maternal age, pre-pregnancy BMI, gestational diabetes mellitus (GDM), preeclampsia, caesarean delivery, gestational age, and neonatal gender. Each potential confounder was added in the model one at a time, and the confounder was kept in the model if the changes of estimates were greater than 10% [28]. As a result, the association of GWG with LGA and SGA was adjusted for maternal age, pregnancy BMI, GDM, and preeclampsia. As there was an interaction effect between GWG and pre-pregnancy BMI on adverse birth weights (P < 0.05), subgroup analyses stratified by pre-pregnancy BMI category were performed in the same manner as above described.

Since the traditional cut-off age for AMA was 35 years [9], participants were further categorized into two maternal age groups: 31-34 years (n 12,189) and ≥ 35 years (n 7665), in order to examine the age-specific association between GWG and adverse birth weights. Logistic regression stratifying by maternal age was performed, and models for LGA and SGA were adjusted for pregnancy BMI, GDM, and preeclampsia.

To test the robustness of the results, the same analysis was repeated by using GWG rate in the second and third trimesters, and the adequacy of GWG rates was defined based on the 2009 IOM recommendation [23]. All analyses were performed using SAS 9.3. P<0.05 was considered as statistically significant.

A post-hoc power test (SAS proc power procedure) showed that the power was >0.999 for the impact of abnormal GWG on adverse birth weight (*n* 19,854), and was >0.999 when the analysis was stratified by maternal age group (*n* 12,189 for 31–34 years and *n* 7665 for \geq 35 years). The power was higher than the criteria (0.8) suggested by Cohen [29], indicating that the number of participants enrolled was sufficient.

Results

General characteristics of the study population

The general characteristics of the total participants and characteristics according to maternal age are listed in Table 1. A total of 19,854 women were included in this study with an average age of 34.0 years and gained 12.5 kg over pregnancy. 9.9% of women were classified as underweight and 27.5% were overweight/obese before pregnancy. After stratified by maternal age, women aged 35 years and older were more likely to be overweight/ obese before pregnancy, but gained less weight during pregnancy than those in the 31-34 years age group (all P < 0.0001). Compared with the 31–34 years age group, the prevalence of GDM, preeclampsia, caesarean delivery, and LGA were significantly higher in the \geq 35 years age group (all P < 0.05). There were no significant differences in the prevalence of SGA between two maternal age groups.

Maternal and neonatal characteristics according to GWG category are presented in Table S1. 27.4% of the women gained weight below the IOM recommendation and 31.2% gained weight above the recommendation. Women with excessive weight gain were more likely to develop GDM and preeclampsia, and deliver LGA infants, whereas women with inadequate weight gain were more likely to deliver SGA infants (all P < 0.0001).

Table 1 General characteristics of the study participants according to maternal age

| Characteristics | Total | Maternal age | P-value | |
|--|-------------------|-------------------|-------------------|----------|
| | (n=19,854) | 31–34 yr | ≥35 yr | |
| | | (n = 12,189) | (n=7665) | |
| Mothers | | | | |
| Pre-pregnancy BMI (kg/cm ²) ^a | 21.3 (19.7, 23.2) | 21.1 (19.5, 22.9) | 21.7 (20.1, 23.5) | < 0.0001 |
| Underweight (n (%)) | 1965 (9.9) | 1422 (11.7) | 543 (7.1) | < 0.0001 |
| Normal weight (n (%)) | 12,477 (62.8) | 7812 (64.1) | 4665 (60.9) | |
| Overweight/obese (n (%)) | 5412 (27.3) | 2955 (24.2) | 2457 (32.1) | |
| Gestational weight gain (kg) | 12.5 (10.0, 15.0) | 13.0 (10.0, 15.0) | 12.0 (9.5, 15.0) | < 0.0001 |
| Rate of second and third trimester weight gain (kg/wk) | 0.46 (0.36, 0.55) | 0.47 (0.38, 0.56) | 0.44 (0.34, 0.54) | < 0.0001 |
| Gestational diabetes mellitus (n (%)) | 5193 (26.2) | 2800 (23.0) | 2393 (31.2) | < 0.0001 |
| Preeclampsia (n (%)) | 349 (1.8) | 180 (1.5) | 169 (2.2) | < 0.0001 |
| Caesarean delivery (n (%)) | 13,723 (69.1) | 7754 (63.6) | 5969 (77.9) | < 0.0001 |
| Newborns | | | | |
| Gestational age at delivery (weeks) | 39.1 (38.7, 39.7) | 39.3 (38.7, 39.9) | 39.0 (38.5, 39.4) | < 0.0001 |
| Gender, females (n (%)) | 9640 (48.6) | 5889 (48.3) | 3751 (48.9) | 0.39 |
| Birth length (cm) | 50.0 (49.0, 51.0) | 50.0 (49.0, 51.0) | 50.0 (48.0, 51.0) | 0.0002 |
| Birth weight (g) | 3300 (3035, 3570) | 3300 (3040, 3570) | 3300 (3030, 3570) | 0.38 |
| Preterm birth (n (%)) | 881 (4.4) | 490 (4.0) | 391 (5.1) | 0.0003 |
| Macrosomia (n (%)) | 926 (4.7) | 575 (4.7) | 351 (4.6) | 0.65 |
| Large for gestational age (n (%)) | 1592 (8.0) | 936 (7.7) | 656 (8.6) | 0.03 |
| Small for gestational age (n (%)) | 1133 (5.7) | 725 (6.0) | 408 (5.3) | 0.06 |

Data are expressed as median (25%, 75%) or n (%)

^a Pre-pregnancy BMI was based on WHO Asians [22]

Table 2 Association between gestational weight gain and adverse birth weights according to the IOM guidelines in total cohort

| GWG | LGA | | | SGA | | |
|-----------------------------|-------------------------------|-------------------|-------------------------|------------|----------------------|-------------------------|
| | n (%) | Crude OR (95% CI) | Adjusted OR (95% CI) | n (%) | Crude OR (95% Cl) | Adjusted OR (95% CI) |
| Overall ^a (n = 1 | 9,854) | | | | | |
| Below | 183 (3.4) | 0.47 (0.39, 0.55) | 0.48 (0.40, 0.57) ** | 472 (8.7) | 1.67 (1.46, 1.91) | 1.63 (1.42, 1.86) ** |
| Within | 571 (7.0) | 1.00 | 1.00 | 441 (5.4) | 1.00 | 1.00 |
| Above | 838 (13.5) | 2.09 (1.87, 2.34) | 1.77 (1.57, 1.99) ** | 220 (3.6) | 0.65 (0.55, 0.76) | 0.72 (0.60, 0.85) * |
| Underweight | ^b (n=1965) | | | | | |
| Below | 5 (0.7) | 0.20 (0.07, 0.46) | 0.19 (0.07, 0.46) * | 110 (15.3) | 2.08 (1.53, 2.83) | 2.02 (1.48, 2.76) ** |
| Within | 34 (3.4) | 1.00 | 1.00 | 79 (8.0) | 1.00 | 1.00 |
| Above | 19 (7.5) | 2.26 (1.25, 4.00) | 2.32 (1.28, 4.11) * | 19 (7.5) | 0.93 (0.54, 1.53) | 0.92 (0.53, 1.51) |
| Normal weigh | nt ^b (n = 12,477) | | | | | |
| Below | 152 (3.6) | 0.51 (0.42, 0.62) | 0.49 (0.41, 0.60) ** | 316 (7.6) | 1.44 (1.22, 1.70) | 1.44 (1.22, 1.71) ** |
| Within | 376 (6.9) | 1.00 | 1.00 | 293 (5.4) | 1.00 | 1.00 |
| Above | 332 (11.7) | 1.78 (1.52, 2.08) | 1.81 (1.55, 2.11) ** | 98 (3.4) | 0.63 (0.49, 0.79) | 0.61 (0.48, 0.76) ** |
| Overweight/0 | Obese ^b (n = 5412) | | | | | |
| Below | 26 (4.7) | 0.50 (0.32, 0.74) | 0.47 (0.30, 0.71) * | 46 (8.4) | 2.24 (1.52, 3.29) | 2.32 (1.55, 3.44) ** |
| Within | 161 (9.1) | 1.00 | 1.00 | 69 (3.9) | 1.00 | 1.00 |
| Above | 487 (15.8) | 1.87 (1.55, 2.26) | 1.94 (1.60, 2.36) ** | 103 (3.3) | 0.85 (0.62, 1.16) | 0.76 (0.56, 1.05) |

^a Adjusted for maternal age, pre-pregnancy BMI, gestational diabetes mellitus, and preeclampsia

^b Adjusted for maternal age, gestational diabetes mellitus, and preeclampsia

*P<0.05, **P<0.0001

Abbreviations: IOM, Institute of Medicine; GWG, gestational weight gain; LGA, large-for-gestational-age; SGA, small-for-gestational-age

The association between GWG and birth weights

The association of GWG with LGA and SGA for the whole study population is shown in Table 2. Compared to women who gained weight within the IOM

recommendation, those gained excessive weight had higher risks of delivering LGA infants (OR 1.77, 95% CI 1.57–1.99, P<0.0001), and lower risk of delivering SGA infants (OR 0.72, 95% CI 0.60–0.85, P<0.05). Conversely,

women who gained insufficient weight had higher risks of delivering SGA infants (OR 1.63, 95% CI 1.42–1.86, P<0.0001), and lower risk of delivering LGA infants (OR 0.48, 95% CI 0.40–0.57, P<0.0001). The associations for LGA remained significant after the women were categorized into different pre-pregnancy BMI groups, while the protective effect of excessive GWG on SGA was only observed in normal weight women.

a) Overall



P<0.05

31-34 yr

(n=1422)

P<0.05

ns

≥35 yr

(n=543)

Below IOM

c) Underweight

P<0.05

7.0

ີ ເ ົີ 6.0[.]

Odds ratio (62% 0.0 0.0 0.0 0.0

1.0 0.0

years: OR 1.58, 95% CI 1.31–1.90, P<0.0001). After the women were categorized into different pre-pregnancy BMI groups, the association remained significant in the 31–34 years age group, whereas for \geq 35 years age group, the association between excessive GWG and the risk of

weights

b) Normal weight

Maternal age-specific association between GWG and birth

The assessment of GWG on the risks of LGA and SGA

was further performed in two maternal age groups

(Figs. 1 and 2). In both maternal age groups, excessive

weight gain was associated with an increased risk of LGA

(31–34 years: OR 1.90, 95% CI 1.63–2.21, *P*<0.0001; ≥35







Fig. 2 Age-specific association between gestational weight gain and small-for-gestational-age according to the IOM guidelines a) Models were adjusted for pre-pregnancy BMI, gestational diabetes mellitus, and preeclampsia; **b-d**) Models were adjusted for gestational diabetes mellitus and preeclampsia

LGA was only significant in normal weight and overweight/obese women. Moreover, among the overweight/ obese women, the effect of excessive GWG on the risk of LGA was more pronounced in the younger age group (31–34 years: OR 2.18, 95% CI 1.68–2.88, P<0.0001; ≥35 years: OR 1.71, 95% CI 1.30–2.25, P<0.001).

Women who gained weight below the IOM recommendation was associated with a higher risk of SGA in both maternal age groups (31–34 years: OR 1.72, 95% CI 1.45–2.04, P<0.0001; ≥35 years: OR 1.46, 95% CI 1.16–1.83, P<0.05). After stratified by pre-pregnancy BMI, while inadequate GWG was still a risk factor for SGA in all BMI categories in the 31–34 years age group, for ≥35 years age group, the significant association was only observed in women who were underweight and overweight/obese. Moreover, among the overweight/ obese women, the magnitude of the association between inadequate GWG and higher risk of SGA was greater for older age group (31–34 years: OR 2.08, 95% CI 1.19–3.55, P<0.05; ≥35 years: OR 2.65, 95% CI 1.47–4.74, P<0.05).

Association between GWG rate in the second and third trimesters and birth weights

The rate of GWG in the second and third trimesters was further estimated. The proportions of women with inadequate, adequate, and excessive GWG rate were 17.6%, 33.5%, and 48.9%, respectively. Results for GWG rate were generally consistent with total GWG regarding LGA and SGA (Table S2 and Table S3), suggesting that stronger associations between abnormal GWG and adverse birth weights were mainly observed in women aged 31–34 years.

Discussion

The present study demonstrated that among overweight and obese women, the effect of insufficient GWG on the risk of SGA was greater with increasing maternal age, whereas the association between excessive GWG and the risk of LGA was more evident in those aged 31-34years. Moreover, the effect of abnormal GWG on the risk of LGA and SGA was not observed in underweight and normal weight women aged ≥ 35 years, respectively.

Prospective cohort studies of GWG and adverse birth weights in older Chinese women have been scarce. The present study was conducted in this older population (31–34 years and \geq 35 years), confirming the association between excessive GWG and higher risk of LGA, but also showing that the effect was more pronounced in women aged 31-34 years, especially for those who were underweight and overweight/obese before pregnancy. These results are likely in part attributable to the variations in the amount of weight gained during pregnancy between two maternal age groups, given that women who aged 31–34 years had more GWG than those aged \geq 35 years. A higher proportion of overweight and obese women in the \geq 35 years age group than in the 31–34 years age group may partially account for the difference in GWG between two groups. Overweight and obese women usually receive greater attention from gynecologists and obstetricians due to their higher risks of adverse perinatal outcomes [17, 30, 31] and are recommended to gain less weight during pregnancy [23].

The present study did not include women aged ≤ 30 years, however, it is noticed that the risk estimate of excessive GWG on LGA in women aged 31-34 years of the present study was higher than most of studies that conducted in Chinese women with mean maternal age at 25-30 years (OR 1.90 vs. OR 1.42-1.70) [32-34], even though their GWG was greater. Considering that some studies found that the risks of LGA were similar in women aged 31–34 years and \geq 35 years [13, 15], it is speculated that maternal age had a higher effect than GWG on the risk of LGA when comparisons were made between women aged 25-30 and 31-34 years, whereas GWG had more influence on LGA when comparisons were made between 31 and 34 years and \geq 35 years. While further research with a wide maternal age category is warranted to verify these results and to explore the underlying mechanisms, the present study emphasizes the need to increase attention to women with childbearing age at 31-34 years, who are becoming increasingly Page 7 of 9

prevalent in China but have been neglected since the commonly used definition of AMA is \geq 35 years.

Unlike LGA, the effect of insufficient GWG on the risk of SGA was more evident for those aged \geq 35 years among overweight and obese women in this study. Given that the risk estimate of insufficient GWG on SGA in prior research that conducted in overweight/obese Chinese women with mean maternal age < 30 years was smaller than that of the present study (OR 1.18-1.40 vs. OR 2.32) [31, 35], it seems like that the effect of insufficient GWG on the risk of SGA in overweight/obese women was greater with increasing maternal age. This observation, however, was not seen in other pre-pregnancy BMI categories. Overweight and obesity is a complex metabolic state, and evidence has indicated that overweight/ obesity is not only associated with fetal overgrowth but also increases the risk of SGA [2] despite of others have reported inconsistent results [30, 36]. The impaired placental function and abnormal transfer of nutrients through placenta in overweight and obese women was regarded as one of the possible mechanisms [37]. Considering that older maternal age is also closely related with placental defects [38], it is likely that some connections may exist between older maternal age and maternal overweight/obesity which exacerbate the effect of insufficient GWG on SGA, but further investigation is required to confirm such a possibility.

It is worth noting that the adverse effect of insufficient GWG on the risk of SGA was not observed in normal weight women aged \geq 35 years. Although the reasons behind it are uncertain, this finding suggests that both maternal age and pre-pregnancy BMI influence the association between GWG and adverse birth weights. The current data again point out the necessity to focus more on women aged 31–34 years, given that the stronger associations between abnormal GWG and adverse birth weights were mainly observed in this age group.

Overall, the present study showed that older women with excessive or insufficient GWG do not appear to be at the same risk for adverse birth weights compared to their younger counterparts. Recommending a single optimal GWG range for women of different maternal ages might be inappropriate. To date, very limited study [39] has been conducted to investigate the optimal weight gain for women with different maternal ages. The current findings may stimulate future investigation to improve the current GWG guidelines by considering different characteristics (e.g. maternal age) of women. In addition, this study highlights the need for healthcare professionals to increase concerns about excessive GWG and LGA in women aged 31-34 years, given that this population is increasing in China and infants born with LGA are closely associated with childhood obesity that has become a serious health problem in the world.

This was the first study to investigate the associations between GWG and adverse birth weights stratified by pre-pregnancy BMI in older Chinese women, and the results were then compared in women aged 31–34 years and those of AMA. The findings may contribute to the limited information available on this population by highlighting the importance of GWG as well as pre-pregnancy BMI on neonatal growth and development. The large sample size, prospective design, and rigorous collection of data allowed a valid assessment of GWG on the risks of adverse birth weights. Moreover, additional measure of GWG rate in the second and third trimesters was performed to control the effect of the length of gestation on adverse birth weights.

Limitations of the present study should be noted. Prepregnancy weight was self-reported which might result in recall error, however, some evidence has indicated that self-reported and clinically measured pre-pregnancy weights were highly correlated [40]. While the number of overall study participants was large, the sample size in the subpopulation such as underweight women, overweight or obese women aged \geq 35 years was limited. Because of this reason, the overweight and obese women were analyzed together. Further research with a large sample size in different maternal ages and pre-pregnancy BMIs is required to assess the association between GWG and adverse birth weights.

Conclusions

In conclusion, the present study demonstrated that maternal age might influence the association between GWG and adverse birth weights in Chinese older women. More attention should be paid to those aged 31–34 years, who have been neglected but their risks of adverse perinatal outcomes were as high as in those of AMA. More importantly, the stronger associations between abnormal GWG and adverse birth weights were mainly seen in the age group of 31–34 years. Future research on association of GWG with other perinatal outcomes such as GDM, preeclampsia and preterm birth in women of different maternal ages are needed to confirm this conclusion, and the results will help to determine whether women of different maternal ages warrant separate GWG recommendations.

Abbreviations

- AMA Advanced maternal age
- BMI Body mass index
- CI Confidence intervals
- GDM Gestational diabetes mellitus
- GWG Gestational weight gain
- IOM Institute of Medicine
- LGA Large-for-gestational-age
- OR Odds ratio
- SGA Small-for-gestational-age

Supplementary Information

The online version contains supplementary material available at https://doi. org/10.1186/s12884-023-06231-y.

Supplementary Material 1

Acknowledgements

All participants and their families in this study are gratefully acknowledged. The authors also thank all colleagues and students working in the study for their continuous and valuable help.

Author contributions

GC designed research; YG and XW conducted research; YW and SS analyzed data; YW wrote paper; YX, GC and BZ revised paper; GC provided critical revisions to the manuscript. All authors have read and approved the final manuscript.

Funding

The study was supported by National Natural Science Foundation of China (81673158) and Chengdu Science and Technology Bureau (2019-GH02-00058-HZ).

Data availability

The datasets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

Declarations

Ethics approval and consent of participate

The study was approved by the Ethics Committee of Sichuan University. All participants provided written informed consent for all study content as well as for linkage of their data from the Medical Birth Registry. All methods were performed in accordance with the relevant guidelines and regulations.

Consent for publication

Not applicable.

Competing interests

The authors declare no competing interests.

Author details

¹Department of Epidemiology and Biostatistics, Institute of Systems Epidemiology, and West China-PUMC C. C. Chen Institute of Health, West China School of Public Health and West China Fourth Hospital, Sichuan University, Chengdu, China

²Department of Gynaecology, West China Women's and Children's Hospital, Sichuan University, Chengdu, China

³Laboratory of Molecular Translational Medicine, Center for Translational Medicine, Key Laboratory of Birth Defects and Related Diseases of Women and Children (Sichuan University), Department of Pediatrics, West China Women's and Children's Hospital, Sichuan University, Chengdu, China

Received: 20 October 2023 / Accepted: 27 December 2023 Published online: 05 January 2024

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